Notes on Bell’s Theorem∗

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1 Bell's Theorem

In this first section, we reconstruct several versions of Bell’s theorem. We work with a standard set-up – to be described in class – involving a pair of photons in the singlet state, and consider the probability that they will both pass through polarizer sheets having specified orientations.

We take $pr_{QM}(A, B|a, b)$ to be the probability, given by QM, for a joint outcome of $A$ on the left and $B$ on the right, given polarizer orientations $a$ on the left and $b$ on the right. (Thus $A$ and $B$ take as values “yes” (the photon passes through the polarizer) and “no” (it does not pass through the polarizer), while $a$ and $b$ take as values lines in the plane orthogonal to the motion of the particles.) We take $pr_{QM}(A, -|a, b)$ to be the probability of outcome $A$ on the left (regardless of the outcome on the right) given polarizer settings $a$ on the left and $b$ on the right. Since the outcome on the right must be either “yes” or “no” (and cannot be both) we have, for example,

$$pr_{QM}(yes, -|a, b) = pr_{QM}(yes, yes|a, b) + pr_{QM}(yes, no|a, b).$$

(Of course, $pr_{QM}(-, B|a, b)$ is handled similarly.)

The predictions of QM are fully characterized by the following two conditions. (Here $\angle(a, b)$ is the (acute) angle between $a$ and $b$.)

(QM1) For all $a, b$, $pr_{QM}(yes, yes|a, b) = \frac{1}{2} \cos^2 \angle(a, b)$.

(QM2) (“yes-no” symmetry) For all $a, b$,

$$pr_{QM}(yes, no|a, b) = pr_{QM}(yes, yes|a, b^\perp)$$
$$pr_{QM}(no, yes|a, b) = pr_{QM}(yes, yes|a^\perp, b)$$
$$pr_{QM}(no, no|a, b) = pr_{QM}(yes, yes|a^\perp, b^\perp).$$

(Here, $b^\perp$ is understood to be the line orthogonal to $b$ (in the plane orthogonal to the motion of the particles).) In light of the symmetry conditions in (QM2), we lose nothing in what follows if we restrict consideration to “yes-yes” outcomes.) It follows from (QM1) and (QM2) that the “single side” probabilities generated by QM satisfy the following condition.

(QM3) For all $a, b$, $pr_{QM}(yes, -|a, b) = \frac{1}{2} = pr_{QM}(-, yes|a, b)$.

The computation is straight-forward. Since $\angle(a, b^\perp) = \frac{\pi}{2} - \angle(a, b)$, we have

$$pr_{QM}(yes, -|a, b) = pr_{QM}(yes, yes|a, b) + pr_{QM}(yes, no|a, b)$$
$$= pr_{QM}(yes, yes|a, b) + pr_{QM}(yes, yes|a, b^\perp)$$
$$= \frac{1}{2} \cos^2 \angle(a, b) + \frac{1}{2} \cos^2 \left(\frac{\pi}{2} - \angle(a, b)\right)$$
$$= \frac{1}{2} \cos^2 \angle(a, b) + \frac{1}{2} \sin^2 \angle(a, b) = \frac{1}{2}. $$
(The other case, of course, is handled similarly.) (QM3) asserts that the (single side) probability that
a photon will pass through a polarizer is $\frac{1}{2}$, whatever the orientation of the polarizer, even though the
joint probability for passage through both polarizers is a function of $\angle(a, b)$. Thus, except for the very
special case in which $\angle(a, b) = \frac{\pi}{4}$ (and hence $\frac{1}{2} \cos^2(\angle(a, b)) = \frac{1}{4}$),
$$ \Pr_{QM}(\text{yes, yes}|a, b) \neq \Pr_{QM}(\text{yes, -}|a, b) \cdot \Pr_{QM}(-, \text{yes}|a, b), $$
i.e., the outcomes on the two sides are statistically correlated.

The “EPR” case, where $b = a^\perp$ (and so $\angle(a, b) = \frac{\pi}{2}$), is of special interest. We have the following
conditions.

(QM4) For all $c$,
$$ \Pr_{QM}(\text{yes, yes}|c, c^\perp) = 0 = \Pr_{QM}(\text{no, no}|c, c^\perp) $$
$$ \Pr_{QM}(\text{yes, no}|c, c^\perp) = \frac{1}{2} = \Pr_{QM}(\text{no, yes}|c, c^\perp). $$

(The two equalities in the first line follow immediately from (QM1) and (QM2). The two in the second
line follow from the first two and (QM3). So, for example,
$$ \Pr_{QM}(\text{yes, no}|c, c^\perp) = \Pr_{QM}(\text{yes, -}|c, c^\perp) - \Pr_{QM}(\text{yes, yes}|c, c^\perp) = \frac{1}{2} - 0 = \frac{1}{2}. $$
Notice that in the EPR case, the outcomes on the two sides exhibit perfect anti-correlation, i.e., the
probability that they differ (either in the pattern “yes-no” or “no-yes”) is 1.

........................................

Now we consider possible “hidden-variable theories” that posit a space $\Lambda$ of hidden states, and deter-
mine probabilities $\Pr_{HV}(A, B|a, b; \lambda)$ for each $\lambda \in \Lambda$. We show that if $\Pr_{HV}$ satisfies certain constraints,
then it is not possible to recover $\Pr_{QM}$ from $\Pr_{HV}$ – more precisely, it is not possible to represent $\Pr_{QM}$
in the form
$$ \Pr_{QM}(A, B|a, b) = \int_{\Lambda} \Pr_{HV}(A, B|a, b; \lambda) \rho(\lambda) d\lambda $$
where $\rho$ is a probability density on $\Lambda$, i.e., an (integrable) function $\rho : \Lambda \to [0, 1]$ satisfying $\int_{\Lambda} \rho(\lambda) d\lambda = 1.$
(We think of $\rho(\lambda)$ as the probability that the photon pair is in hidden state $\lambda$.)

Technical note: We have been deliberately vague in the preceding paragraph. We have made reference
to “spaces” and to integration over those spaces without indicating exactly what mathematical structures
and operations we have in mind. We have done so because it really makes no difference. The proofs that
follow invoke only such elementary properties of integrals as are exhibited by all species. We could be
thinking about Riemann integration over suitably chosen sets in $\mathbb{R}^n$, such as one studies in a calculus
course. Or, for example, we could be thinking about integration over abstract measure spaces.

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We will be interested in the following four conditions on hidden variable theories.

**Quasi-determinism:** For all $A, B, a, b, \lambda$, $\text{pr}_{HV}(A, B|a, b; \lambda) = 0/1$

**Screening-off:** For all $A, B, a, b, \lambda$, $\text{pr}_{HV}(A, B|a, b; \lambda) = \text{pr}_{HV}(A, \neg|a, b; \lambda) \cdot \text{pr}_{HV}(\neg, B|a, b; \lambda)$

**Locality:** For all $A, B, a, a', b, b', \lambda$,

\[
\text{pr}_{HV}(A, \neg|a, b; \lambda) = \text{pr}_{HV}(A, \neg|a, b'; \lambda)
\]

\[
\text{pr}_{HV}(\neg, B|a, b; \lambda) = \text{pr}_{HV}(\neg, B|a', b; \lambda)
\]

**Anti-correlation:** For all $A, c, \lambda$, $\text{pr}_{HV}(A, A|c, \perp; \lambda) = 0$.

The first is clear. It asserts that conditionalization on the hidden state $\lambda$ pushes all HV-probabilities to 0 or 1. (Notice that our notion of hidden variable theory does not build-in this condition from the beginning, i.e., we allow for “non-deterministic hidden variable theories”.) The second asserts that the observed correlation between outcomes (yes or no) on the two sides is “screened-off” by the underlying hidden variable state. The third asserts that (after conditionalization on the hidden state $\lambda$) the HV-probability for an outcome on one side is independent of the polarizer orientation (or setting) on the other side. (The distinction between “outcome-outcome correlations” and “setting-outcome correlations” is crucially important here.) The fourth condition asserts that HV-probabilities (after conditionalization on the hidden state $\lambda$) exhibit the same the anti-correlation pattern as QM-probabilities. (Recall condition (QM4).)

We turn to our first two versions of Bell’s theorem. (There will be three altogether.) The first rules out hidden variable theories that satisfy the locality and screening-off conditions. The second (really just a corollary of the first) rules out theories that satisfy the locality and quasi-determinism conditions.

**Proposition 1.1.** Let $\text{pr}_{HV}(A, B|a, b; \lambda)$ satisfy the locality and screening-off conditions. Let $\rho$ be a probability density on $\Lambda$, and let $\text{pr}_{HV}(A, B|a, b)$ be defined by

\[
\text{pr}_{HV}(A, B|a, b) = \int_{\Lambda} \text{pr}_{HV}(A, B|a, b; \lambda) \rho(\lambda) d\lambda.
\]

Then it is not the case that $\text{pr}_{QM}(A, B|a, b) = \text{pr}_{HV}(A, B|a, b)$ for all $A, B, a, b$.

To prove the theorem, we show that if the stated hypotheses hold, then $\text{pr}_{HV}$ must satisfy the following inequality (the “Clauser-Horne inequality”):

\[
0 \leq \text{pr}_{HV}(\text{yes}, \neg|a, \neg) + \text{pr}_{HV}(\neg, \text{yes}|
eg, b) + \text{pr}_{HV}(\text{yes}, \text{yes}|a', b') - \text{pr}_{HV}(\text{yes}, \text{yes}|a, b') - \text{pr}_{HV}(\text{yes}, \text{yes}|a', b) \leq 1.
\]

for all $a, b, a', b'$. (The expressions $\text{pr}_{HV}(\text{yes}, \neg|a, \neg)$ and $\text{pr}_{HV}(\neg, \text{yes}|
eg, b)$ make sense if the locality condition holds for then

\[
\text{pr}_{HV}(\text{yes}, \neg|a, b) = \int_{\Lambda} \text{pr}_{HV}(\text{yes}, \neg|a, b; \lambda) \rho(\lambda) d\lambda
\]
does not depend on \(b\) (and similarly \(pr_{HV}(\_, \text{yes}|a, b)\) does not depend on \(a\)). This will suffice, since QM probabilities do not satisfy the counterpart inequality for all \(a, b, a', b'\). For example, if \(\angle(a', b') = \angle(a, b) = \pi/2\) and \(\angle(a, b) = \angle(a', b') = \angle(a', b) = \pi/6\), then

\[
pr_{QM}(\text{yes, yes}|a, b) = pr_{QM}(\text{yes, yes}|a', b) = \frac{1}{2} \cos^2\left(\frac{\pi}{6}\right) = \frac{3}{8}
\]

\[
pr_{QM}(\text{yes, yes}|a', b') = \frac{1}{2} \cos^2\left(\frac{\pi}{2}\right) = 0
\]

\[
pr_{QM}(\_, \text{yes}|a, \_) = pr_{QM}(\_, \text{yes}|\_, b) = \frac{1}{2}
\]

(The expressions \(pr_{QM}(\_, \text{yes}|a, \_)\) and \(pr_{QM}(\_, \text{yes}|\_, b)\) make sense since, by (QM4), \(pr_{QM}(\_, \text{yes}|a, b)\) does not depend on \(b\), and \(pr_{QM}(\_, \text{yes}|a, b)\) does not depend on \(a\).) Hence the sum of six terms in the inequality is

\[
\frac{1}{2} + \frac{1}{2} + 0 - 3\left(\frac{3}{8}\right) = -\frac{1}{8}
\]

(which is not between 0 and 1).

**Proof.** First note that for all numbers \(x, x', y, y'\) in the interval \([0, 1]\),

\[
0 \leq x + y + x'y' - xy' - x'y - xy \leq 1. \tag{1.1}
\]

There are various ways to see this. One involves a simple consideration of three cases: (i) \(x \leq x'\), (ii) \(y \leq y'\), (iii) \(x > x'\) and \(y > y'\). In case (i), we have

\[
0 \leq x(1 - y) + y(1 - x') + y'(x' - x) \leq x + (1 - x') + (x' - x) = 1.
\]

But the underlined expression is equal to \(x + y + x'y' - xy' - x'y - xy\). So (*) holds. Similarly, in case (ii), we have

\[
0 \leq y(1 - x) + x(1 - y') + x'(y' - y) \leq y + (1 - y') + (y' - y) = 1.
\]

Finally, in case (iii), we have

\[
0 \leq (x - x')(y - y') + y(1 - x) + x(1 - y) \leq xy + y(1 - x) + (1 - y) = 1.
\]

Thus (1.1) holds in all three cases. Now assume \(pr_{HV}\) satisfies the locality and screening-off conditions. Then we can express \(pr_{HV}(\text{yes, yes}|a, b; \lambda)\) in the form

\[
pr_{HV}(\text{yes, yes}|a, b; \lambda) = pr_{HV}(\_, a; \_, \text{yes}\_|\_, \_, \lambda) \cdot pr_{HV}(\_, \text{yes}|\_, a; \_, \_; \lambda).
\tag{1.2}
\]

Given lines \(a, a', b, b'\), and hidden state \(\lambda\), let

\[
x = pr_{HV}(\_, \_, \_, a; \_, \_, \lambda)
\]
\[
x' = pr_{HV}(\_, \_, a; \_, \_, \text{yes}; \_, \_, \lambda)
\]
\[
y = pr_{HV}(\_, \_, \_, \_, b; \_, \_, \lambda)
\]
\[
y' = pr_{HV}(\_, \text{yes}|\_, \_, \_, b'; \_, \_, \lambda).
\]
Then $x, x', y, y'$ are all in the interval $[0,1]$ and so, by (1.1) and (1.2),
\[
0 \leq \text{pr}_{HV}(\text{yes}, \underline{a}, \underline{\_}; \lambda) + \text{pr}_{HV}(\underline{\_}, \text{yes}|\underline{b}; \lambda) + \text{pr}_{HV}(\text{yes}, \text{yes}|a', b'; \lambda) - \text{pr}_{HV}(\text{yes}, \text{yes}|a, b'; \lambda) - \text{pr}_{HV}(\text{yes}, \text{yes}|a', b; \lambda) - \text{pr}_{HV}(\text{yes}, \text{yes}|a, b; \lambda) \leq 1.
\]
We move from this inequality to the Clauser-Horne inequality with a simple integration. Let $X(a, a', b, b', \lambda)$ be the (middle) sum of six terms, and let $\rho$ be a probability density on $\Lambda$. Then
\[
0 = \int_\Lambda 0 \cdot \rho(\lambda) d\lambda \leq \int_\Lambda X(a, a', b, b', \lambda) \cdot \rho(\lambda) d\lambda \leq \int_\Lambda 1 \cdot \rho(\lambda) d\lambda = 1,
\]
and
\[
\int_\Lambda X(a, a', b, b', \lambda) \cdot \rho(\lambda) d\lambda = \text{pr}_{HV}(\text{yes}, \underline{a}, \underline{\_}; \lambda) + \text{pr}_{HV}(\underline{\_}, \text{yes}|\underline{b}; \lambda) + \text{pr}_{HV}(\text{yes}, \text{yes}|a', b'); \lambda) - \text{pr}_{HV}(\text{yes}, \text{yes}|a, b'; \lambda) - \text{pr}_{HV}(\text{yes}, \text{yes}|a', b; \lambda) - \text{pr}_{HV}(\text{yes}, \text{yes}|a, b; \lambda).
\]

It is not hard to show that the quasi-determinism condition implies the screening off condition (and we leave this as an exercise).

**Problem 1.1.** Show

(a) quasi-determinism $\Rightarrow$ screening-off

(b) screening-off $&$ locality $&$ anti-correlation $\Rightarrow$ quasi-determinism

So we have the following immediate corollary. (Everything remains the same except that reference to the latter condition is replaced by reference to the former.)

**Proposition 1.2.** Let $\text{pr}_{HV}(A, B|a, b; \lambda)$ satisfy the locality and quasi-determinism conditions. Let $\rho$ be a probability density on $\Lambda$, and let $\text{pr}_{HV}(A, B|a, b)$ be defined by
\[
\text{pr}_{HV}(A, B|a, b) = \int_\Lambda \text{pr}_{HV}(A, B|a, b; \lambda) \rho(\lambda) d\lambda.
\]
Then it is not the case that
\[
\text{pr}_{QM}(A, B|a, b) = \text{pr}_{HV}(A, B|a, b)
\]
for all $A, B, a, b$.

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Now we consider a third version of the theorem that is a bit different in character from the first two. It is of interest in its own right. (For one thing, it is perfectly precise in formulation. No reference is
made to some not fully specified sense of integration.) And it will prepare the way for our discussion of Itamar Pitowsky’s work on the geometric interpretation of Bell type inequalities.

For the moment, let us continue within the framework of section 1.2. Suppose we have a space Λ of hidden states for our two-photon system, a probability function \( pr_{HV}(\text{yes, yes}|a, b; \lambda) \) over Λ, and a probability density over Λ, i.e., an (integrable) function \( \rho : \Lambda \rightarrow [0, 1] \) satisfying \( \int_{\Lambda} \rho(\lambda)d\lambda = 1. \) Further suppose that \( pr_{HV}(A, B|a, b) \) is defined (as above) by

\[
pr_{HV}(A, B|a, b) = \int_{\Lambda} pr_{HV}(A, B|a, b; \lambda)\rho(\lambda)d\lambda.
\]

Finally, suppose that \( pr_{HV}(A, B|a, b; \lambda) \) satisfies both locality and quasi-determinism.

For all lines \( a \) and \( b, \) consider sets:

\[
X_{ab} = \{ \lambda \in \Lambda : pr_{HV}(\text{yes, yes}|a, b; \lambda) = 1 \}
\]

\[
L_{a} = \{ \lambda \in \Lambda : pr_{HV}(\text{yes}|a, -; \lambda) = 1 \}
\]

\[
R_{b} = \{ \lambda \in \Lambda : pr_{HV}(-, \text{yes}|-, b; \lambda) = 1 \}
\]

The latter two are well defined by locality. By locality and quasi-determinism (which implies screening off)

\[
pr_{HV}(\text{yes, yes}|a, b; \lambda) = pr_{HV}(\text{yes, -}|a, -; \lambda) \cdot pr_{HV}(-, \text{yes}|-, b; \lambda).
\]

So

\[
pr_{HV}(\text{yes, yes}|a, b; \lambda) = 1 \iff pr_{HV}(\text{yes, -}|a, -; \lambda) = 1 \text{ and } pr_{HV}(-, \text{yes}|-, b; \lambda) = 1.
\]

Hence,

\[
\lambda \in X_{ab} \iff \lambda \in L_{a} \land \lambda \in R_{b}
\]

or, equivalently,

\[
X_{ab} = L_{a} \cap R_{b}
\]

for all \( a \) and \( b. \)

Now consider the measure \( \mu \) on \( \Lambda \) defined by setting

\[
\mu(C) = \int_{C} \rho(\lambda)d\lambda.
\]

(We understand \( C \) to be in the domain of \( \mu \) if the integral is well defined.) It then follows by quasi-determinism that, for all \( a \) and \( b, \)

\[
pr_{HV}(\text{yes, yes}|a, b) = \int_{\Lambda} pr_{HV}(\text{yes, yes}|a, b; \lambda)\rho(\lambda)d\lambda
\]

\[
= \int_{X_{ab}} 1 \cdot \rho(\lambda)d\lambda + \int_{(\Lambda - X_{ab})} 0 \cdot \rho(\lambda)d\lambda = \int_{X_{ab}} \rho(\lambda)d\lambda
\]

\[
= \mu(X_{ab}) = \mu(L_{a} \cap R_{b}).
\]
(Notice that it is quasi-determinism that allows us to divide the first integral into two subintegrals – one
over the set $X_{ab}$ where $pr_{HV}(yes,yes|a,b;\lambda)$ is 1, and one over the complement set ($X - X_{ab}$) where
$pr_{HV}(yes,yes|a,b;\lambda)$ is 0.) Similarly, it follows that

$$
pr_{HV}(yes,\neg|a,b) = \int_{\Lambda} pr_{HV}(yes,\neg|a,\neg;\lambda)\rho(\lambda)d\lambda = \int_{L_a} \rho(\lambda)d\lambda = \mu(L_a)
$$

$$
pr_{HV}(\neg,yes|a,b) = \int_{\Lambda} pr_{HV}(\neg,yes|\neg,b;\lambda)\rho(\lambda)d\lambda = \int_{R_b} \rho(\lambda)d\lambda = \mu(R_b)
$$

for all $a$ and $b$. Now suppose it were the case that $pr_{QM}(A,B|a,b) = pr_{HV}(A,B|a,b)$ for all $a, b$. Then
it would follow that

$$
\begin{align*}
pr_{QM}(yes,yes|a,b) & = \mu(L_a \cap R_b) \\
pr_{QM}(yes,\neg|a,\neg) & = \mu(L_a) \\
pr_{QM}(\neg,yes|\neg,b) & = \mu(R_b)
\end{align*} 
$$

for all $a$ and $b$.

We now have another way to set up the second version of the theorem. We forget about our route to
the three equations in (*) and use them, in effect, to characterize a (local, deterministic) hidden variable
theory. We think of $\mu$ as just some probability measure (or other) on $\Lambda$, and pay no attention to whether
it arises from a probability density. For any subset $C$ of $X$ (in the domain of $\mu$), we understand $\mu(C)$ as
the probability that the exact underlying state of the system is, in fact, in $C$. In particular, we interpret
$\mu(L_a \cap R_b)$ as the probability that the underlying state of the system happens to be in one in which it
is determined that, with orientations $a$ and $b$, both photons will pass through the polarizer sheets. The
surprising result, of course, is that we cannot have a hidden variable theory in the sense just characterized.
This will be our third version of Bell’s theorem.

Before stating it, we need to recall the definition of a probability space. It is a structure $(X, \Sigma, \mu)$ where
$X$ is a non-empty set, $\Sigma$ is a set of subsets of $X$ satisfying three conditions

(F1) $X \in \Sigma$

(F2) For all subsets $A$ of $X$, $A \in \Sigma \Rightarrow (X - A) \in \Sigma$

(F3) For all subsets $A_1$, $A_2$ of $X$, $A_1, A_2 \in \Sigma \Rightarrow (A_1 \cup A_2) \in \Sigma$

and $\mu$ is a probability measure on $\Sigma$, i.e., a map $\mu : \Sigma \rightarrow [0,1]$ such that

(M1) $\mu(X) = 1$

(M2) For all sets $A_1$, $A_2$ in $\Sigma$ that are pairwise disjoint, $\mu(A_1 \cup A_2) = \mu(A_1) + \mu(A_2)$.

It follows immediately by induction, of course, that the conditions in (F3) and (M2) can be extended
to arbitrary finite unions. Furthermore, it follows from (F2) and (F3) that $\Sigma$ is closed under finite
intersections (as well as unions) – since

$$(A_1 \cap A_2 \cap \ldots \cap A_n) = X - [(X - A_1) \cup (X - A_2) \cup \ldots \cup (X - A_n)].$$
Technical note: Standardly one replaces (F3) with the stronger requirement that $\Sigma$ be closed under countable unions, and replaces (M2) with the stronger requirement of countable additivity. But the difference between the two formulations (finite versus countable) is irrelevant for our purposes.

Now we have all the pieces in place.

**Proposition 1.3.** There does not exist a classical probability space $(X, \Sigma, \mu)$ and, for all directions $a, b$, sets $L_a$ and $R_b$ in $\Sigma$, such that the three equations in (*) hold.

(There are clear advantages to this formulation. The disadvantage is that it depends crucially on the assumption of quasi-determinism. There is no variant that uses only the screening-off condition (and so serves as counterpart to our first version of Bell’s thorem).)

**Proof.** We already know that quantum mechanical probabilities violate the Clauser-Horne inequality. So it suffices to show that given a probability space $(X, \Sigma, \mu)$, and any four sets $L_a, L'_a, R_b, R'_b$ in $\Sigma$, the corresponding inequality

$$0 \leq \mu(L_a) + \mu(R_b) + \mu(L'_a \cap R'_b) - \mu(L_a \cap R_b) - \mu(L'_a \cap R_b) - \mu(L_a \cap R'_b) \leq 1$$

is satisfied. We do so using a low brow computation. Every point $\lambda$ in $X$ is either in $L_a$ or in its complement $L_a^c = X - L_a$. Similarly, it is either in $L'_a$ or in its complement $L'_a^c$; the same for $R_b$ and $R_b^c$. Thus we can partition $X$ into sixteen disjoint sets. Typical members are

$$(L_a \cap L'_a \cap R_b^c \cap R'_b^c)$$
$$(L_a^c \cap L'_a \cap R_b^c \cap R'_b^c)$$

The first is the set of all points $\lambda$ in $X$ that are in $L_a$, $L'_a$, and $R_b$, but not in $R_b$. The second is the set of all such points that are in $R_b$, but not in $L_a$, $L'_a$, or $R_b$.

To proceed, we just have to express the six sets that enter into the Clauser-Horne inequality in terms of the elements of the partition, and then do some cancelling. For example, $(L_a \cap R_b)$ can be expressed as the disjoint union of the four sets

$$(L_a \cap L'_a \cap R_b \cap R'_b)$$
$$(L_a \cap L'_a \cap R_b \cap R'_b^c)$$
$$(L_a \cap L'_a \cap R_b^c \cap R'_b)$$
$$(L_a \cap L'_a \cap R_b^c \cap R'_b^c)$$

(Notice that $L_a$ and $R_b$ appear in all four (rather than their complements $L_a^c$ and $R_b^c$).) So it follows
We list the decompositions for the six sets below. For ease of reading, we use a simple labelling scheme for the 16 sets in the partition. So, for example, \( \mu(1011) \) stands for the set \( \mu(L_a \cap L_{a'} \cap R_b \cap R_{b'}) \). (The marks on the extreme right (checks, circles, infinity signs) indicate a cancellation pattern that we will consider shortly.)
\[
\begin{align*}
\mu(L_a \cap R_b) &= \mu(L_a \cap L_a' \cap R_b \cap R_b') \quad \mu(1111) \circ \\
+\mu(L_a \cap L_a' \cap R_b' \cap R_b') &\quad \mu(1101) \circ \circ \\
+\mu(L_a \cap L_a' \cap R_b \cap R_b') &\quad \mu(1011) \circ \circ \circ \\
+\mu(L_a \cap L_a' \cap R_b' \cap R_b') &\quad \mu(1001) \circ \circ \circ \circ \\
\mu(L_a' \cap R_b) &= \mu(L_a \cap L_a' \cap R_b \cap R_b') \quad \mu(1111) \circ \circ \\
+\mu(L_a \cap L_a' \cap R_b \cap R_b') &\quad \mu(1101) \circ \circ \circ \\
+\mu(L_a^- \cap L_a' \cap R_b \cap R_b') &\quad \mu(0111) \circ \circ \circ \circ \\
+\mu(L_a^- \cap L_a' \cap R_b \cap R_b') &\quad \mu(0101) \circ \circ \circ \circ \circ \\
\mu(L_a \cap R_b) &= \mu(L_a \cap L_a' \cap R_b \cap R_b') \quad \mu(1111) \circ \circ \circ \circ \\
+\mu(L_a \cap L_a' \cap R_b \cap R_b') &\quad \mu(1101) \circ \circ \circ \circ \\
+\mu(L_a^- \cap L_a' \cap R_b \cap R_b') &\quad \mu(0111) \circ \circ \circ \circ \circ \\
+\mu(L_a^- \cap L_a' \cap R_b \cap R_b') &\quad \mu(0101) \circ \circ \circ \circ \circ \circ \\
\end{align*}
\]

To compute the Clauser-Horne expression

\[
\mu(L_a) + \mu(R_b) + \mu(L_a' \cap R_b') - \mu(L_a \cap R_b') - \mu(L_a' \cap R_b) - \mu(L_a \cap R_b)
\]

we add the \((8+8+4)\) terms arising from

\[
\mu(L_a) + \mu(R_b) + \mu(L_a' \cap R_b')
\]

and then subtract the \((4+4+4)\) terms arising from

\[
\mu(L_a \cap R_b') + \mu(L_a' \cap R_b) + \mu(L_a \cap R_b)
\]

But as a simple inspection confirms, every term in the second group already appears in the first group, and eight terms in the first group are left over. (See the cancellation pattern indicated with check marks and related symbols.) When the dust clears, we have

\[
\mu(L_a) + \mu(R_b) + \mu(L_a' \cap R_b') - \mu(L_a \cap R_b') - \mu(L_a' \cap R_b) - \mu(L_a \cap R_b)
\]

\[
=\mu(1100) + \mu(1000) + \mu(1010) + \mu(0111) + \mu(0011) + \mu(0010) + \mu(1101) + \mu(0101).
\]

The right side is a sum of terms, each of which is \(\geq 0\). So the sum is \(\geq 0\). And the sum is clearly \(\leq 1\) since the sum of all 16 terms of form \(\mu(\quad \quad \quad)\) is 1. So the expression on the left side of the equality is clearly bounded by 0 and 1, as claimed.
2 The Geometric Interpretation of Bell Type Inequalities

Let us put quantum mechanics aside for the moment and consider a quite general question that Pitowsky [5] poses and answers.

Let \( n \geq 2 \) be given, and let \( S \) be a non-empty subset of \( \{ \langle i, j \rangle : 1 \leq i < j \leq n \} \). Further, assume we are given \( n + |S| \) numbers

\[
p_i \quad i = 1, \ldots, n
\]
\[
p_{ij} \quad \langle i, j \rangle \in S
\]

(Here \( |S| \) is the number of elements in \( S \).) It is helpful to think of the numbers as determining an \((n + |S|)-\)tuple \( \langle p_1, \ldots, p_n, \ldots, p_{ij}, \ldots \rangle \) where, let us agree, the \( p_{ij} \) are ordered by their indices, and the latter are ordered lexicographically. We say that the \((n + |S|)-\)tuple admits a probability space representation if there exists a probability space \((X, \Sigma, \mu)\), and (not necessarily distinct) sets \( A_1, \ldots, A_n \in \Sigma \) such that, for all \( i \in \{1, 2, \ldots, n\} \) and all \( \langle i, j \rangle \in S \),

\[
p_i = \mu(A_i)
\]
\[
p_{ij} = \mu(A_i \cap A_j).
\]

**Question:** Under what conditions does \( \langle p_1, \ldots, p_n, \ldots, p_{ij}, \ldots \rangle \) admit a probability space representation?

Pitowsky gives a beautifully simple answer. Let \( \{0, 1\}^n \) be the set of all \( n \)-tuples of 0’s and 1’s. Given any such \( n \)-tuple \( \epsilon = \langle \epsilon_1, \ldots, \epsilon_n \rangle \), let \( p^\epsilon \) be the \((n + |S|)-\)tuple \( \langle \epsilon_1, \ldots, \epsilon_n, \ldots, \epsilon_i \epsilon_j, \ldots \rangle \) where the product term \( \epsilon_i \epsilon_j \) appears precisely if \( \langle i, j \rangle \in S \). (For example, if \( n = 3, S = \{\langle 1, 2 \rangle, \langle 2, 3 \rangle\} \), and \( \epsilon = \langle 0, 1, 1 \rangle \),

\[
p^\epsilon = \langle 0, 1, 1, 0, 1 \rangle.
\]

Now let \( c(n, S) \) to be the closed, convex polytope in \( \mathbb{R}^{(n + |S|)} \) whose vertices are the \( 2^n \) vectors of form \( p^\epsilon \), where \( \epsilon \in \{0, 1\}^n \), i.e., the set of all vectors that can be expressed as convex sums of these \( 2^n \) vectors. (Recall that, quite generally, given vectors \( v_1, \ldots, v_m \) in a vector space (over \( \mathbb{R} \)), a convex sum of those vectors is a sum of the form \( \lambda_1 v_1 + \ldots + \lambda_m v_m \) where \( \lambda_i \geq 0 \) for all \( i \), and \( \lambda_1 + \ldots + \lambda_m = 1 \).)

**Example** If \( n = 2 \) and \( S = \{\langle 1, 2 \rangle\} \), \( c(n, S) \) is the set of all vectors in \( \mathbb{R}^3 \) of the form

\[
\lambda(0, 0) (0, 0, 0) + \lambda(1, 0) (1, 0, 0) + \lambda(0, 1) (0, 1, 0) + \lambda(1, 1) (1, 1, 1)
\]
\[
= ( \lambda(1, 0) + \lambda(1, 1)), \ (\lambda(0, 1) + \lambda(1, 1)), \ \lambda(1, 1) )
\]

where the four coefficients \( \lambda(0, 0), \lambda(1, 0), \lambda(0, 1), \lambda(1, 1) \) are non-negative and sum to 1. (See figure 2.1.)

**Proposition 2.1.** For all \( n \) and \( S \), \( \langle p_1, \ldots, p_n, \ldots, p_{ij}, \ldots \rangle \) admits a probability space representation iff it belongs to \( c(n, S) \).
Proof. Let $n$, $S$, and $p$ be given. Assume first that $p$ admits a classical representation, i.e., assume there is a probability space $(X, \Sigma, \mu)$ and sets $A_1, \ldots, A_n \in \Sigma$ such that, for all $i \leq n$ and all $\langle i, j \rangle \in S$, $p_i = \mu(A_i)$ and $p_{ij} = \mu(A_i \cap A_j)$. Given a set $A \in \Sigma$, let $A^1 = A$ and $A^0 = X - A$. Further, given $\epsilon = \langle \epsilon_1, \ldots, \epsilon_n \rangle \in \{0, 1\}^n$, let

$$A(\epsilon) = A_{\epsilon_1} \cap A_{\epsilon_2} \cap \ldots \cap A_{\epsilon_n}.$$ 

The sets $A(\epsilon)$ form a partition of $X$ as $\epsilon$ ranges over $\{0, 1\}^n$, i.e., $\epsilon \neq \epsilon' \Rightarrow A(\epsilon) \cap A(\epsilon') = \emptyset$ and $\cup\{A(\epsilon) : \epsilon \in \{0, 1\}^n\} = X$. Finally, let $\lambda(\epsilon) = \mu(A(\epsilon))$. Clearly, $\lambda(\epsilon) \geq 0$ for all $\epsilon$ in $\{0, 1\}^n$, and $\Sigma_{\epsilon \in I} \lambda(\epsilon) = 1$, where $I = \{0, 1\}^n$. For all $i \in \{1, 2, \ldots, n\}$ and all $\langle i, j \rangle \in S$,

$$A_i = \cup \{A(\epsilon) : \epsilon \in I \text{ and } \epsilon_i = 1\}$$
$$A_i \cap A_j = \cup \{A(\epsilon) : \epsilon \in I \text{ and } \epsilon_i = \epsilon_j = 1\}.$$

Hence,

$$p_i = \mu(A_i) = \Sigma_{\epsilon \in I : \epsilon_i = 1} \lambda(\epsilon) = \Sigma_{\epsilon \in I} \lambda(\epsilon) \epsilon_i = \Sigma_{\epsilon \in I} \lambda(\epsilon)(p^\epsilon)_i$$
$$p_{ij} = \mu(A_i \cap A_j) = \Sigma_{\epsilon \in I : \epsilon_i = \epsilon_j = 1} \lambda(\epsilon) = \Sigma_{\epsilon \in I} \lambda(\epsilon) \epsilon_i \epsilon_j = \Sigma_{\epsilon \in I} \lambda(\epsilon)(p^\epsilon)_{ij}.$$ 

Thus $p = \Sigma_{\epsilon \in I} \lambda(\epsilon)p^\epsilon$. So $p$ belongs to $c(n, S)$.

Conversely, assume $p$ belongs to $c(n, S)$. Then there exist numbers $\lambda(\epsilon) \geq 0$ such that $\Sigma_{\epsilon \in I} \lambda(\epsilon) = 1$ and $p = \Sigma_{\epsilon \in I} \lambda(\epsilon)p^\epsilon$. Let

$$X = I = \{0, 1\}^n$$
$$\Sigma = \mathcal{P}(X) \text{ (the power set of } X)$$

Figure 2.1: The $c(n, S)$ polytope in the case where $n = 2$ and $S = \{(1, 2)\}$. The planes of the bounding faces are identified.
and for all $A$ in $\Sigma$, let

$$\mu(A) = \sum_{\epsilon \in A} \lambda(\epsilon).$$

Clearly, $(X, \Sigma, \mu)$ is a probability space. Finally, for all $i \leq n$, let

$$A_i = \{ \epsilon : \epsilon_i = 1 \}.$$

Then

$$\mu(A_i) = \sum_{\epsilon \in A_i} \lambda(\epsilon) = \sum_{\epsilon \in I : \epsilon_i = 1} \lambda(\epsilon) = \sum_{\epsilon \in I} \lambda(\epsilon) \epsilon_i = \mu;$$

$$\mu(A_i \cap A_j) = \sum_{\epsilon \in A_i \cap A_j} \lambda(\epsilon) = \sum_{\epsilon \in I : \epsilon_i = \epsilon_j = 1} \lambda(\epsilon) = \sum_{\epsilon \in I} \lambda(\epsilon) \epsilon_i \epsilon_j$$

$$= \sum_{\epsilon \in I} \lambda(\epsilon)(p')_{ij} = p_{ij}.$$

So $p$ admits a classical representation. \hfill \Box

The polytope $c(n, S)$ can be characterized not only as the convex hull of its vertices, but also as the set of vectors bounded by its supporting hyperplanes, and thus as the set of vectors whose components satisfy a particular set of linear inequalities. It turns out that it is precisely these hyperplane-describing inequalities, for simple choices of $n$ and $S$, that we have come to know as "Bell-type inequalities".

**Proposition 2.2.** (Examples)

(a) Let $n = 2$ and let $S = \{ (1, 2) \}$. A vector $(p_1, p_2, p_{12})$ belongs to $c(n, S)$ in this case iff

$$0 \leq p_{12} \leq p_1 \leq 1$$

$$0 \leq p_{12} \leq p_2 \leq 1$$

$$p_1 + p_2 - p_{12} \leq 1.$$

(b) Let $n = 3$ and let $S = \{ (1, 2), (1, 3), (2, 3) \}$. A vector $(p_1, p_2, p_3, p_{12}, p_{13}, p_{23})$ belongs to $c(n, S)$ in this case iff for all $\langle i, j \rangle \in S$,

$$0 \leq p_{ij} \leq p_i \leq 1$$

$$0 \leq p_{ij} \leq p_j \leq 1$$

$$p_i + p_j - p_{ij} \leq 1$$

$$p_1 + p_2 + p_3 - p_{12} - p_{13} - p_{23} \leq 1$$

$$0 \leq p_1 - p_{12} - p_{13} + p_{23}$$

$$0 \leq p_2 - p_{12} - p_{23} + p_{13}$$

$$0 \leq p_3 - p_{13} - p_{23} + p_{12}.$$
(c) Let $n = 4$ and let $S = \{(1,3), (1,4), (2,3), (2,4)\}$. A vector $\langle p_1, p_2, p_3, p_4, p_{13}, p_{14}, p_{23}, p_{24} \rangle$ belongs to $c(n, S)$ in this case iff for all $\langle i, j \rangle \in S$,

\[
\begin{align*}
0 &\leq p_{ij} \leq p_i \leq 1 \\
0 &\leq p_{ij} \leq p_j \leq 1 \\
p_i + p_j - p_{ij} &\leq 1 \\
-1 &\leq p_{13} + p_{14} + p_{24} - p_{23} - p_1 - p_4 \leq 0 \\
-1 &\leq p_{23} + p_{24} + p_{14} - p_{13} - p_2 - p_4 \leq 0 \\
-1 &\leq p_{14} + p_{13} + p_{23} - p_{24} - p_1 - p_3 \leq 0 \\
-1 &\leq p_{24} + p_{23} + p_{13} - p_{14} - p_2 - p_3 \leq 0.
\end{align*}
\]

Of course, if one combines propositions 2.1 and 2.2, one can bypass reference to correlation polytopes, and assert directly that, in the particular cases considered, vectors in $\mathbb{R}^{n+|S|}$ admit a classical representation if and only if they satisfy the corresponding set of inequalities. The combined versions, for cases (b) and (c), were proved by Fine ([2], [3]). His argument, however, did not involve geometrical ideas, and did not readily lend itself to generalization. Pitowsky’s does. Proposition 2.1 provides an algorithm for finding the set of “generalized Bell inequalities” corresponding to any choice of $n$ and $S$. One can find it, at least in principle, by systematically formulating all conditions of form “vector $p$ falls to one (specified) side of hyperplane $H$”, as $H$ ranges over all bounding hyperplanes of $c(n, S)$.

Proof. (a) Let $n = 2$ and let $S = \{(1,2)\}$. The vertices of $c(n, S)$ are: $(0,0,0), (1,0,0), (0,1,0)$, and $(1,1,1)$. All four satisfy the inequalities:

\[
0 \leq p_{12} \leq p_1 \leq 1 \quad 0 \leq p_{12} \leq p_2 \leq 1 \quad p_1 + p_2 - p_{12} \leq 1.
\]

(This is easy to check.) Furthermore, if vectors $p$ and $p'$ in $\mathbb{R}^{n+|S|}$ satisfy the inequalities, so does every convex combination $p'' = \lambda p + (1 - \lambda) p'$. For example, $p''$ satisfies the third inequality because

\[
p''_{12} + p''_2 - p''_{12} = [\lambda p + (1 - \lambda) p']_1 + [\lambda p + (1 - \lambda) p']_2 - [\lambda p + (1 - \lambda) p']_{12} \\
= \lambda [p_1 + p_2 - p_{12}] + (1 - \lambda) [p'_1 + p'_2 - p'_{12}] \leq \lambda + (1 - \lambda) = 1.
\]

Thus every vector in $c(n, S)$, i.e., every convex combination of the four vertices, satisfies the three inequalities.

Conversely, assume that the vector $p$ in $\mathbb{R}^{n+|S|}$ satisfies them. Then we can express $p$ as a convex combination of $(0,0,0), (1,0,0), (0,1,0)$, and $(1,1,1)$:

\[
p = (p_1, p_2, p_{12}) \\
= (1 - p_1 - p_2 + p_{12})(0,0,0) + (p_1 - p_{12})(1,0,0) + (p_2 - p_{12})(0,1,0) + p_{12}(1,1,1).
\]
That is, \( p \) belongs to \( c(n, S) \).

(b) Let \( n = 3 \) and let \( S = \{\langle 1, 2\rangle, \langle 1, 3\rangle, \langle 2, 3\rangle\} \). To prove that every vector in \( c(n, S) \) satisfies the indicated inequalities it suffices to check that every vertex

\[
p^\epsilon = (\epsilon_1, \epsilon_2, \epsilon_3, \epsilon_1\epsilon_2, \epsilon_1\epsilon_3, \epsilon_2\epsilon_3, \epsilon_2\epsilon_4) \quad \epsilon = \langle \epsilon_1, \epsilon_2, \epsilon_3 \rangle \in \{0, 1\}^3
\]

does so. (Again, this is easy.) For the converse, assume that a vector \( p \) in \( \mathbb{R}^{n+|S|} \) satisfies the inequalities. Then there exists a number \( \alpha \) such that

\[
\alpha \leq \min\{p_{12}, p_{13}, p_{23}, 1 - (p_1 + p_2 + p_3 - p_{12} - p_{13} - p_{23})\}
\]

and

\[
\max\{0, (-p_1 + p_{12} + p_{13}), (-p_2 + p_{12} + p_{23}), (-p_3 + p_{13} + p_{23})\} \leq \alpha.
\]

(The inequalities guarantee that every number in the second list is less than or equal to every number in the first list.) To every \( \epsilon \) in \( \{0, 1\}^3 \) we assign a number \( \lambda(\epsilon) \geq 0 \) as follows:

\[
\begin{align*}
\lambda(0,0,0) &= 1 - (p_1 + p_2 + p_3 - p_{12} - p_{13} - p_{23}) - \alpha \\
\lambda(1,0,0) &= \alpha + (p_1 - p_{12} - p_{13}) \\
\lambda(0,1,0) &= \alpha + (p_2 - p_{12} - p_{23}) \\
\lambda(0,0,1) &= \alpha + (p_3 - p_{13} - p_{23}) \\
\lambda(1,1,0) &= p_{12} - \alpha \\
\lambda(1,0,1) &= p_{13} - \alpha \\
\lambda(0,1,1) &= p_{23} - \alpha \\
\lambda(1,1,1) &= \alpha.
\end{align*}
\]

Clearly the sum of these eight numbers is 1. Furthermore, \( p = \Sigma_{\epsilon \in \{0,1\}^3} \lambda(\epsilon)p^\epsilon \). For example,

\[
\lambda(1,0,0) + \lambda(1,1,0) + \lambda(1,0,1) + \lambda(1,1,1) = [\alpha + (p_1 - p_{12} - p_{13})] + (p_{12} - \alpha) + (p_{13} - \alpha) + \alpha = p_1
\]

and

\[
\lambda(1,1,0) + \lambda(1,1,1) = (p_{12} - \alpha) + \alpha = p_{12}
\]

as required. (The other cases are handled similarly.)

(c) Let \( n = 4 \) and let \( S = \{\langle 1, 3\rangle, \langle 1, 4\rangle, \langle 2, 3\rangle, \langle 2, 4\rangle\} \). To prove that every vector in \( c(n, S) \) satisfies the indicated inequalities it suffices to check that every vertex

\[
p^\epsilon = (\epsilon_1, \epsilon_2, \epsilon_3, \epsilon_4, \epsilon_1\epsilon_3, \epsilon_1\epsilon_4, \epsilon_2\epsilon_3, \epsilon_2\epsilon_4) \quad \epsilon = \langle \epsilon_1, \epsilon_2, \epsilon_3, \epsilon_4 \rangle \in \{0, 1\}^4
\]

does so. For the converse, assume that a vector \( p \) in \( \mathbb{R}^{n+|S|} \) satisfies the inequalities. Then there exists a number \( \beta \) such that

\[
\beta \leq \min\{p_1, p_2, (p_1 - p_{13} + p_{23}), (p_2 - p_{23} + p_{13}), (p_1 - p_{14} + p_{24}), (p_2 - p_{24} + p_{14})\}
\]
\[
\max\{0, (p_1 + p_2 - 1), (p_{13} + p_{23} - 3), (p_{14} + p_{24} - 4),
(p_1 + p_2 + 3 - p_{13} - p_{23} - 1), (p_1 + p_2 + 4 - p_{14} - p_{24} - 1)\} \leq \beta.
\]

(Again, the inequalities guarantee that every number in the second list is less than or equal to every number in the first list.) Now let \(S' = \{(1, 3), (2, 3), (1, 3)\}\), and consider \(p' = (p'_1, p'_2, p'_3, p'_4, p'_5, p'_6)\) in \(\mathbb{R}^{3+|S'|}\) defined by
\[
\begin{align*}
p'_1 &= p_1 & p'_2 &= p_2 & p'_3 &= p_3 \\
p'_4 &= \beta & p'_5 &= p_4 & p'_6 &= p_{23}
\end{align*}
\]
One can easily check that \(p'\) satisfies all the inequalities cited in part (b). For example,
\[
p'_1 + p'_2 + p'_3 - p'_4 - p'_5 - p'_6 \leq 1
\]
holds since the left side expression equals \((p_1 + p_2 + 3 - \beta - p_{13} - p_{23})\) and \((p_1 + p_2 + 3 - p_{13} - p_{23} - 1) \leq \beta\).

Hence, \(p'\) belongs to \(c(n, S')\), i.e., \(p'\) can be expressed as a convex sum of form
\[
p' = \sum_{\epsilon \in \{0, 1\}^3} \lambda'(\epsilon) p'.
\]

Similarly, the vector \(p'' = (p''_1, p''_2, p''_3, p''_4, p''_5, p''_6)\) in \(\mathbb{R}^{3+|S'|}\) defined by
\[
\begin{align*}
p''_1 &= p_1 & p''_2 &= p_2 & p''_3 &= p_3 \\
p''_4 &= \beta & p''_5 &= p_4 & p''_6 &= p_{23}
\end{align*}
\]
can be expressed as the convex sum
\[
p'' = \sum_{\epsilon \in \{0, 1\}^3} \lambda''(\epsilon) p'.
\]

Now for \(\epsilon = (\epsilon_1, \epsilon_2, \epsilon_3, \epsilon_4) \in \{0, 1\}^4\) we set
\[
\lambda(\epsilon) = \lambda(\epsilon_1, \epsilon_2, \epsilon_3, \epsilon_4) = \frac{\lambda'(\epsilon_1, \epsilon_2, \epsilon_3) \lambda''(\epsilon_1, \epsilon_2, \epsilon_4)}{\lambda'(\epsilon_1, \epsilon_2, 0) + \lambda'(\epsilon_1, \epsilon_2, 1)}
\]
if the denominator is not zero. If it is zero, we set \(\lambda(\epsilon) = \lambda(\epsilon_1, \epsilon_2, \epsilon_3, \epsilon_4) = 0\).

(Note that we could just as well have taken \(\lambda''(\epsilon_1, \epsilon_2, 0) + \lambda''(\epsilon_1, \epsilon_2, 1)\) for the denominator because the two expressions are equal for all \(\epsilon_1\) and \(\epsilon_2\). This follows from the fact that

(i) \(\lambda'(1, 1, 0) + \lambda'(1, 1, 1) = \lambda'_{12} = \lambda''(1, 1, 0) + \lambda''(1, 1, 1)\)

(ii) \(\lambda'(0, 1, 0) + \lambda'(0, 1, 1) + \lambda'(1, 1, 0) + \lambda'(1, 1, 1) = \lambda''(0, 1, 0) + \lambda''(1, 0, 0) + \lambda''(1, 0, 1) + \lambda''(1, 1, 1)\)

(iii) \(\lambda'(0, 1, 0) + \lambda'(0, 1, 1) + \lambda'(1, 1, 0) + \lambda'(1, 1, 1) = \lambda''(0, 1, 0) + \lambda''(0, 1, 1) + \lambda''(1, 1, 0) + \lambda''(1, 1, 1)\)
There are four cases to consider. If $\epsilon_1 = 1$ and $\epsilon_2 = 1$, the desired equation is (i). If $\epsilon_1 = 1$ and $\epsilon_2 = 0$, we derive it by subtracting (i) from (ii). Similarly, if $\epsilon_1 = 0$ and $\epsilon_2 = 1$, we derive it by subtracting (i) from (iii). Finally, if $\epsilon_1 = 0$ and $\epsilon_2 = 0$, the desired equation follows from the three previous ones and the fact that

$$\Sigma_{\epsilon \in \{0,1\}^3} \lambda'(\epsilon) = 1 = \Sigma_{\epsilon \in \{0,1\}^3} \lambda''(\epsilon).$$

It is clear that $\lambda(\epsilon) \geq 0$ for all $\epsilon \in \{0,1\}^4$. We also have $\Sigma_{\epsilon \in \{0,1\}^3} \lambda(\epsilon) = 1$. To see this note first that

$$\Sigma_{\epsilon \in \{0,1\}^3} \lambda(\epsilon) = \frac{\Sigma_{\epsilon_1, \epsilon_2, \epsilon_3} \lambda'(\epsilon_1, \epsilon_2, \epsilon_3) \lambda''(\epsilon_1, \epsilon_2, \epsilon_4)}{\lambda'(\epsilon_1, \epsilon_2, 0) + \lambda'(\epsilon_1, \epsilon_2, 1)}.$$

But for all $\epsilon_1, \epsilon_2$,

$$\Sigma_{\epsilon_3, \epsilon_4} \lambda'(\epsilon_1, \epsilon_2, \epsilon_3) \lambda''(\epsilon_1, \epsilon_2, \epsilon_4) = \lambda'(\epsilon_1, \epsilon_2, 0) + \lambda'(\epsilon_1, \epsilon_2, 1) \lambda''(\epsilon_1, \epsilon_2, 0) + \lambda''(\epsilon_1, \epsilon_2, 1).$$

So

$$\Sigma_{\epsilon \in \{0,1\}^3} \lambda(\epsilon) = \Sigma_{\epsilon_1, \epsilon_2} \lambda''(\epsilon_1, \epsilon_2, 0) + \lambda''(\epsilon_1, \epsilon_2, 1) = \Sigma_{\epsilon \in \{0,1\}^3} \lambda''(\epsilon) = 1.$$

We claim, finally, that $p = \Sigma_{\epsilon \in \{0,1\}^3} \lambda(\epsilon) p'$. We check just one representative component: $p_{14}$. We need to show that

$$p_{14} = \lambda(1, 0, 0, 1) + \lambda(1, 0, 1, 1) + \lambda(1, 1, 0, 1) + \lambda(1, 1, 1, 1).$$

Now if $\lambda'(1, 0, 0) + \lambda'(1, 0, 1) \neq 0$, then

$$\lambda'(1, 0, 0) + \lambda'(1, 0, 1) = \frac{\lambda'(1, 0, 0) \lambda''(1, 0, 1) + \lambda'(1, 0, 1) \lambda''(1, 0, 1)}{\lambda'(1, 0, 0) + \lambda'(1, 0, 1)} = \lambda''(1, 0, 1).$$

On the other hand, if $\lambda'(1, 0, 0) + \lambda'(1, 0, 1) = 0$, then $\lambda(1, 0, 0, 1) + \lambda(1, 0, 1, 1) = 0$. But in this case we also have $\lambda''(1, 0, 0) + \lambda''(1, 0, 1) = 0$, and hence $\lambda''(1, 0, 1) = 0$. So, in either case,

$$\lambda(1, 0, 0, 1) + \lambda(1, 0, 1, 1) = \lambda''(1, 0, 1).$$

Similarly,

$$\lambda(1, 1, 0, 1) + \lambda(1, 1, 1, 1) = \lambda''(1, 1, 1).$$

Therefore,

$$\lambda(1, 0, 0, 1) + \lambda(1, 0, 1, 1) + \lambda(1, 1, 0, 1) + \lambda(1, 1, 1, 1) = \lambda''(1, 0, 1) + \lambda''(1, 1, 1) = p''_{13} = p_{14}$$

as required. □
3 One Attempt to Get Around Bell’s Theorem

Here we consider one recent, somewhat non-standard, response to Bell’s theorem by László E. Szabó. He argues that, the theorem notwithstanding, quantum mechanics is compatible with both “local determinism” and the classical character of probability. (See Szabó ([6], [7]), and Bana and Durt [1].)

Recall the set up in the first section. Given our pair of photons in the singlet state, we know that there exist orientations of the polarizers $a, a’, b, b’$ such that the associated probabilities

\begin{align*}
  p_a &= pr_{QM}(\text{yes}, \neg |a, \neg) \\
  p_b &= pr_{QM}(\neg, \text{yes}|\neg, b) \\
  p_{ab} &= pr_{QM}(\text{yes}, \text{yes}|a, b) \\
  p_{ab'} &= pr_{QM}(\text{yes}, \text{yes}|a, b') \\
  p_{a' b} &= pr_{QM}(\text{yes}, \text{yes}|a', b) \\
  p_{a' b'} &= pr_{QM}(\text{yes}, \text{yes}|a', b')
\end{align*}

have the values

\begin{align*}
  p_a = p_b = \frac{1}{2} & \quad p_{ab} = p_{ab'} = p_{a' b} = \frac{3}{8} \quad p_{a' b'} = 0.
\end{align*}

These violate the Clauser-Horne inequality

$$0 \leq p_a + p_b - p_{ab} - p_{ab'} - p_{a' b} + p_{a' b'} \leq 1.$$ 

Hence (by theorem 1.3.1), we know these “probabilities” do not admit a probability space representation, i.e., there does not exist a probability space $(X, \Sigma, \mu)$ and sets $L_a^+, L_{a'}^+, R_b^+, R_{b'}^+ \in \Sigma$ such that

\begin{align*}
  p_a &= \mu(L_a^+) \\
  p_b &= \mu(R_b^+) \\
  p_{ab} &= \mu(L_a^+ \cap R_b^+) \\
  p_{ab'} &= \mu(L_a^+ \cap R_{b'}^+) \\
  p_{a'b} &= \mu(L_{a'}^+ \cap R_b^+) \\
  p_{a'b'} &= \mu(L_{a'}^+ \cap R_{b'}^+).
\end{align*}

(\*)

One straight-forward interpretation of this result is that “quantum probability” violates the constraints of classical probability (as codified by Kolmogorov). The starting point of Szabó’s response is the observation that the “quantum probabilities” in question here are conditional in character. $p_a$, for example, is supposed to be the conditional probability that the left photon will pass through the polarizer given that the latter is oriented in direction $a$. What if we try to take into consideration just what the probability is that the polarizer is oriented in that direction? Or if we are casting the discussion in terms of determinism, what if we consider possible hidden variables that determine polarizer settings in addition to everything else (rather than treat the settings as independent variables under our control).
Szabó’s proposal, in effect, is to consider a second, weaker sense in which one might try to give the numbers \( p_a, p_b, p_{ab}, \ldots, p_{a' b'} \) a “probability space representation”. Here we explicitly recognize the composite character of the events under consideration. Rather than looking for just four sets \( L^+_a, L^+_a, R^+_b, R^+_b \) in \( \Sigma \), we look for six sets \( L_a, L_a', R_b, R_{b'}, L^+_a, R^+_b \) in \( \Sigma \). Intuitively, we think of \( L_a \) as the set of hidden states in which it is determined that the left polarizer will have orientation \( a \) (and similarly for \( L_a', R_b, R_{b'} \)). We think of \( L^+_a \) as the set of hidden states in which it is determined that the photon will pass through the left polarizer (and similarly for \( R^+_b \)). The conditions we require now are not (\( \ast \ast \)) above, but rather the following:

\[
\begin{align*}
p_a &= \frac{\mu(L_a \cap L^+_a)}{\mu(L_a)} \\
p_b &= \frac{\mu(R_b \cap R^+_b)}{\mu(R_b)} \\
p_{ab} &= \frac{\mu(L_a \cap L^+_a \cap R_b \cap R^+_b)}{\mu(L_a \cap R_b)} \\
p_{ab'} &= \frac{\mu(L_{a'} \cap L^+_a \cap R_b \cap R^+_b)}{\mu(L_{a'} \cap R_b)} \\
p_{a'b} &= \frac{\mu(L_a \cap L^+_a \cap R_{b'} \cap R^+_b)}{\mu(L_a \cap R_{b'})} \\
p_{a'b'} &= \frac{\mu(L_{a'} \cap L^+_a \cap R_{b'} \cap R^+_b)}{\mu(L_{a'} \cap R_{b'})}.
\end{align*}
\]  

(\( \ast \ast \ast \))

Actually, we need more than just these conditions. We are now, implicitly, relativizing our probability space representations (or, equivalently, our deterministic hidden variable theories) to particular experiments. In any one experiment, the polarizer orientations, right and left, occur with particular frequencies. These frequencies must also be recovered. (Maybe on one occasion, for example, the four possibilities \( (a, b), (a, b'), (a', b), (a', b') \) are observed with equal frequency – each, say, occurring 250 times in a run of 1000.) Nothing has been said so far about such frequencies because it has been assumed that they made no difference. Now we imagine that we have a particular experimental run in mind, and have observed experimental probabilities (or frequencies) for the different polarizer settings:

\[
\begin{align*}
l_a &= \text{observed probability for orientation } a \text{ on the left} \\
l_{a'} &= \text{observed probability for orientation } a' \text{ on the left} \\
r_b &= \text{observed probability for orientation } b \text{ on the right} \\
r_{b'} &= \text{observed probability for orientation } b' \text{ on the right}
\end{align*}
\]

What we must add to (\( \ast \ast \ast \)) is the following set of conditions:
\begin{align*}
l_a &= \mu(L_a) \\
l_{a'} &= \mu(L_{a'}) \\
r_b &= \mu(R_b) \\
r_{b'} &= \mu(R_{b'}) \\
l_{a}r_b &= \mu(L_a \cap R_b) \\
l_{a'}r_b &= \mu(L_{a'} \cap R_b) \\
l_{a}r_{b'} &= \mu(L_a \cap R_{b'}) \\
l_{a'}r_{b'} &= \mu(L_{a'} \cap R_{b'})
\end{align*}

Putting all this together, the question under consideration is whether, given a particular run of the two photon experiment, we can find a probability space \((X, \Sigma, \mu)\) and sets \(L_a, L_{a'}, R_b, R_{b'}, L^+, R^+\) in \(\Sigma\) such that (***) and (****) hold. The answer is certainly ‘yes’. Let’s first verify that this is so, and then return to consider the significance of this fact.

<table>
<thead>
<tr>
<th>(R_b)</th>
<th>(L_a)</th>
<th>(L_{a'})</th>
</tr>
</thead>
<tbody>
<tr>
<td>(L_a \cap L^+ \cap R_b \cap R^+)</td>
<td>(\frac{3}{8}(l_ar_b))</td>
<td>(\frac{3}{8}(l_{a'}r_b))</td>
</tr>
<tr>
<td>(L_a \cap L^- \cap R_b \cap R^+)</td>
<td>(\frac{1}{8}(l_ar_b))</td>
<td>(\frac{1}{8}(l_{a'}r_b))</td>
</tr>
<tr>
<td>(L_a \cap L^+ \cap R_{b'} \cap R^+)</td>
<td>(\frac{3}{8}(l_ar_{b'}))</td>
<td>(\frac{3}{8}(l_{a'}r_{b'}))</td>
</tr>
<tr>
<td>(L_a \cap L^- \cap R_{b'} \cap R^+)</td>
<td>(\frac{1}{8}(l_ar_{b'}))</td>
<td>(\frac{1}{8}(l_{a'}r_{b'}))</td>
</tr>
<tr>
<td>(L_{a'} \cap L^+ \cap R_b \cap R^-)</td>
<td>(0)</td>
<td>(\frac{1}{2}(l_{a'}r_b))</td>
</tr>
<tr>
<td>(L_{a'} \cap L^- \cap R_{b'} \cap R^-)</td>
<td>(0)</td>
<td>(\frac{1}{2}(l_{a'}r_{b'}))</td>
</tr>
</tbody>
</table>

Table 1: The displayed probabilities satisfy all conditions in (***) and (****).

It will be easiest to exhibit the requisite example with a diagram (see Table 1). (The elements of the background set \(X\) make no difference. They might as well be points in a region of the Euclidean plane.) In the diagram we label 16 distinct boxes, each the intersection of four sets, e.g., \((L_a \cap L^+ \cap R_b \cap R^+)\). (Notation: \(L^-\) and \(R^-\) are understood to be the complement sets \(X - L^+\) and \(X - R^+\).) The six sets \(L_a, L_{a'}, R_b, R_{b'}, L^+, R^+\) individually, of course, can be reconstructed as appropriate unions of (eight of these) boxes. So, for example, \(L_a\) is the union of the boxes:

\[
(L_a \cap L^+ \cap R_b \cap R^+) \quad (L_a \cap L^+ \cap R_b \cap R^-) \\
(L_a \cap L^- \cap R_b \cap R^+) \quad (L_a \cap L^- \cap R_b \cap R^-) \\
(L_a \cap L^+ \cap R_{b'} \cap R^+) \quad (L_a \cap L^+ \cap R_{b'} \cap R^-) \\
(L_a \cap L^- \cap R_{b'} \cap R^+) \quad (L_a \cap L^- \cap R_{b'} \cap R^-).
\]
In each box there is a displayed a number that should be understood as the probability assigned by the measure $\mu$ to that box. So, for example,

$$\mu(L_a \cap L^+ \cap R_b \cap R^+) = \frac{3}{8}(l_a r_b).$$

Assignments to disjoint unions of these boxes are determined by addition. Thus

$$\mu(L_a \cap R_b) = \mu(L_a \cap L^+ \cap R_b \cap R^+) + \mu(L_a \cap L^- \cap R_b \cap R^-) + \mu(L_a \cap L^- \cap R_b \cap R^+) + \mu(L_a \cap L^+ \cap R_b \cap R^-)$$

$$= \frac{3}{8}(l_a r_b) + \frac{1}{8}(l_a r_b) + \frac{1}{8}(l_a r_b) + \frac{3}{8}(l_a r_b)$$

$$= l_a r_b.$$

It is straightforward to verify that all the conditions in $(* *)$ and $(* * *)$ are satisfied. For example,

$$\frac{\mu(L_a \cap L^+)}{\mu(L_a)} = \frac{(\frac{3}{8} + \frac{1}{8})(l_a r_b) + (\frac{3}{8} + \frac{1}{8})(l_a r_b')}{l_a r_b + l_a r_b'} = \frac{1}{2} = p_a$$

and

$$\frac{\mu(L_a \cap L^+ \cap R_b \cap R^+)}{\mu(L_a \cap R_b)} = \frac{\frac{3}{8}(l_a r_b)}{l_a r_b} = \frac{3}{8} = p_{ab}.$$

The example we have just considered – with quantum mechanical probabilities arising from a pair of photons in the singlet state – is very specific, of course. The question naturally arises whether a similar treatment is available for all probabilities arising in quantum mechanics. The question is not yet precise, and we will not take the time to make it so. But this can be done (see Szabó [7] and Bana and Durt [1]) and the answer is ‘yes’. Roughly speaking, the claim is this.

All probabilities involving experimental trials can be considered conditional in character. They can be understood to be of form $p(O|I)$, the probability that if an experiment characterized by initial conditions $I$ is performed, the outcome will be $O$. Sometimes, as in our example, we consider, side by side, probabilities whose associated initial conditions are incompatible with one another. (We cannot simultaneously test the probability that the left photon will pass through the polarizer when it has orientation $a$, and also test the probability that it will pass through when polarizer has orientation $a'$. ) It is in these cases that it sometimes becomes impossible to give the numbers in question (i.e., the numbers $p(O|I)$) a probability space representation in the initial sense.

The burden of the theorem under consideration is to make precise and prove the claim that in all cases – not just those involving quantum mechanics – one can have a modified probability space representation in which the numbers emerge as conditional probabilities of form $\mu(O \cap I)/\mu(I)$. In this sense, at least, Szabó argues, observed empirical data can never be in conflict with the principles of classical probability or with determinism.
It should be appreciated just how weak this sense is. Let’s stay with the two photon example. It is
a prediction of quantum mechanics that the probability for joint passage through the two polarizers is
given by

$$\text{pr}_{QM}(\text{yes, yes}|a, b) = \frac{1}{2}\cos^2 \angle(a, b).$$

(Recall assertion (QM1) on page 2.) This formula is confirmed by numerous experiments of the most
diverse sort. It seems to express a fact about the the two photon system (in the single state) itself, about
its disposition to behave whenever it is subjected to a test of the appropriate sort. One would like to
have a hidden variable theory that reconstructs these probabilities once and for all, without reference to
particular experimental tests. Instead, one gets from Szabó, in effect, a different hidden variable theory
for each test or, to be more precise, a different theory for each set of non-negative real numbers $l_a, l_{a'},
 r_b, r_{b'}$ summing to 1.

I hope to have further discussion of the significance of Szabo’s work in class.
4 Hidden Variables and “Generalized Probability Spaces”

Here we consider another sense (in addition to the one considered in section 3) in which one can have a hidden variable theory for quantum mechanics, despite the restrictions of Bell’s theorem. It involves a relaxation on the conditions that characterize a (classical) probability space.

Let’s start by considering hidden variables in a general setting, but make use of the notation and terminology introduced in section 2.

Suppose we have a physical system, in a particular state, and a collection of $n$ ($n \geq 2$) experimental tests to which the system can be subjected, each with two possible outcomes (positive and negative). There is no need to be precise about what constitutes a “test”, but it will help to keep in mind an example. We can (once again) think of a photon propagating in the z direction, in a particular state, impinging on a polaroid sheet in the $x - y$ plane, with a particular rotational orientation in that plane, and consider whether the photon is transmitted by the sheet or not. Associated with the $n$ tests are $n$ probabilities $p_1$ through $p_n$ (where, of course, $p_i$ is the probability of a positive outcome in the $i$th test).

We do not take for granted that these tests are pairwise compatible, i.e., that given any two of them, it makes sense to speak of a joint or composite test that is passed precisely if the two are passed individually. But in particular cases it may be possible to do so, and in such cases there will be a well defined joint probability. So in addition to the probabilities $p_1$ through $p_n$, we will have joint probabilities $p_{ij}$ for particular pairs of indices $i$ and $j$ (with $1 \leq i < j \leq n$).

For example, consider two instances of the test involving the photon and the polaroid sheet. In the first, the transmission axis of the sheet is aligned with the x axis. In the second, it is rotated $45^\circ$ away from it (say). The two tests are not compatible in the relevant sense. We might try to perform “both” tests by placing two polaroid sheets in the path of the photon, one behind the other, each with the appropriate orientation, and simply consider the probability that the photon will be transmitted by both sheets. But the resultant probability depends on the order in which the sheets are placed (and it does not make sense to ask which is the “correct” order).

The standard Einstein Podolsky Rosen scenario, in contrast, provides a family of examples of test pairs that are compatible, at least under particular background conditions. Here we consider a system with two subcomponents, “left” and “right”. Test 1 is conducted on the left component, test 2 on the right one, and we require that the tests be performed at points in spacetime (i.e., some points or other, not particular points) that are spacelike related. It appears to be a fact of experimental life that, in such cases, the order in which the tests are performed, as determined relative to any background observer, simply makes no difference. (And if relativity theory is correct, it cannot make a systematic difference, since otherwise the two component system could be used to define an invariant order relation over spacelike related points). So, in these EPR-type cases there is a clear sense in which the two tests can be performed jointly, and there is a well defined (joint) probability for a positive outcome in both.
Once again, let us imagine that we have a physical system, in a particular state, and a collection of \( n \) experimental tests to which the system can be subjected. Let \( S \) be the set of pairs \( \langle i, j \rangle \) such that the \( i \)th and \( j \)th tests are compatible and \( i < j \). (These might arise from an EPR-type scenario, but they need not. We can proceed at a high level of generality.) So there is determined an ordered set of probabilities 

\[
(p_1, \ldots, p_n, \ldots p_{ij}, \ldots)
\]

(ordered as above) for a positive outcome in the individual and joint tests. (We shall call it an \((\text{experimental})\ \text{probability data set of type} \ (n, S)\).)

Now what would it mean to have a “deterministic hidden variable theory” for this data set? Presumably, such a “theory” would allow us to maintain that the outcome to each test (individual and joint) is uniquely determined by the exact underlying state of the system. So there would be a set of exact states \( X \), and a rule of association between individual tests and subsets of \( X \):

\[
\text{\( i \)th test} \rightarrow A_i \subseteq X.
\]

Assigned to the \( i \)th test would be the set of all exact states in which a positive outcome to the \( i \)th test is determined. Further, the theory would allow us to interpret the observed probability \( p_i \) for a positive outcome to the \( i \)th test as but the probability that the system is, in fact, in a state belonging to that associated subset. So there would be a probability function \( \mu \) such that

\[
p_i = \mu(A_i)
\]

for all \( i \). Finally, in the case of a joint test, the theory would allow us to interpret the observed probability for a positive outcome as the probability that the system is in a state belonging to the intersection of the two sets associated individually with the component tests. So we would have

\[
p_{ij} = \mu(A_i \cap A_j)
\]

for all \( \langle i, j \rangle \) in \( S \).

Thus we are led back, quite naturally, to all the elements considered in section 2. Let us say that the probability data set \( (p_1, \ldots, p_n, \ldots p_{ij}, \ldots) \) \textit{admits a hidden variable theory in sense 1} precisely if it admits a probability space representation as characterized in the handout. (There will be a second sense in a moment.) As we know, quantum mechanics generates data sets that do \textit{not} satisfy this condition. In particular, in the Clauser-Horne case

\[
n = 4 \quad S = \{\langle 1, 3 \rangle, \langle 1, 4 \rangle, \langle 2, 3 \rangle, \langle 2, 4 \rangle\}
\]

the data set

\[
(1/2, 1/2, 1/2, 1/2, 3/8, 3/8, 3/8, 0, 3/8)
\]

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does not do so.

Now we come to the heart of the matter. Just what significance does this fact have? The standard view is that it shows that it is not possible to give quantum mechanics a local, deterministic hidden variable interpretation. But there are those who dissent. Arthur Fine, for one, attributes an entirely different significance to it. (See the references listed below.) His point, as I understand it, concerns the definition of a probability space.

Let \((X, \Sigma, \mu)\) be one such. Here \(X\) is a non-empty set. \(\Sigma\) is a set of subsets of \(X\) satisfying three conditions:

\[
\begin{align*}
& (F1) \quad X \in \Sigma \\
& (F2) \quad \text{For all subsets } A \text{ of } X, \ A \in \Sigma \Rightarrow (X - A) \in \Sigma \\
& (F3) \quad \text{For all subsets } A, B \text{ of } X, \ A, B \in \Sigma \Rightarrow (A \cup B) \in \Sigma;
\end{align*}
\]

Finally, \(\mu\) is a probability measure on \(\Sigma\), i.e., a map \(\mu : \Sigma \rightarrow [0, 1]\) such that

\[
\begin{align*}
& (M1) \quad \mu(X) = 1 \\
& (M2) \quad \text{For all sets } A, B \text{ in } \Sigma \text{ that are pairwise disjoint, } \mu(A \cup B) = \mu(A) + \mu(B).
\end{align*}
\]

(Again, we only need to consider finite closure under unions and finite additivity.)

The problem is with \((F3)\). It follows from \((F3)\), in conjunction with \((F2)\), of course, that

\[
(F4) \quad \text{For all sets } A \text{ and } B \text{ in } \Sigma, \ (A \cap B) \in \Sigma.
\]

This condition, in particular, Fine considers inappropriately strong.

Suppose that \((X, \Sigma, \mu)\) is a hidden variable theory in sense 1 for a probabilistic data set \(\langle p_1, ..., p_n, ... p_{ij} ... \rangle\) of type \((n, S)\), and suppose that \(S\) does not include the pair \((k, l)\). That is, assume we are dealing with a case in which the \(k\)th and \(l\)th tests are not compatible. (So it does not make sense to speak of the joint probability that both tests are passed.) Let \(A_k\) and \(A_l\) be the subsets of \(X\) corresponding to the \(k\)th and \(l\)th tests. These sets both belong to \(\Sigma\). So, the intersection \((A_k \cap A_l)\) belongs to \(\Sigma\) as well (by \((F4)\)), and our hidden variable theory assigns to it a probability \(\mu(A_k \cap A_l)\). Thus the theory specifies a probability that the system is in a state in which it is determined that both tests would be passed, even though this joint probability has no empirical significance. Fine argues that we should only require of a hidden variable theory that it return probabilities that are empirically well defined, and that it is not surprising that we run into trouble when we (implicitly) insist that it do more.

“... hidden variables and the Bell inequalities are all about ... imposing requirements to make well defined precisely those probability distributions for noncommuting [i.e., incompatible] observables whose rejection is the very essence of quantum mechanics.” (Fine [3])

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What interests me about Fine’s position is that it suggests a technical question. What happens if we relax the definition of a “probability space” in such a way that (F4) need no longer hold? Then \( \mu(A_k) \) and \( \mu(A_l) \) can be well defined, without it following that \( \mu(A_k \cap A_l) \) is also well defined. There seems to be a natural way to do this. I have in mind using what Stanley Gudder calls a “generalized probability space” (Gudder [4]). This is a structure \((X, \Sigma, \mu)\) where \(X\) is, as before, a non-empty set, \(\Sigma\) is a set of subsets of \(X\) satisfying particular closure conditions, and \(\mu\) is a probability measure on \(\Sigma\), i.e., a map \(\mu : \Sigma \rightarrow [0, 1]\) satisfying conditions (M1) and (M2). The sole difference is that we no longer require that \(\Sigma\) be closed under arbitrary unions. Instead, we require only that \(\Sigma\) be a set of subsets of \(X\) that satisfies (F1), (F2), and, instead of (F3), the weaker condition:

(F3*) For all subsets \(A\) and \(B\) of \(X\) that are pairwise disjoint, \(A, B \in \Sigma \Rightarrow (A \cup B) \in \Sigma\).

Notice that (F3*) is all that is needed for (M2) to be well posed. Notice also that (F3*) does not imply (F4).

Let \(\langle p_1, ..., p_n, ... p_{ij} ... \rangle\) be a probability data set of type \((n, S)\). We shall say that it admits a hidden variable theory in sense 2 if there is a generalized probability space \((X, \Sigma, \mu)\) and (not necessarily distinct) sets \(A_1, ..., A_n \in \Sigma\) such that, for all \(i \in \{1, 2, ..., n\}\) and all \(\langle i, j \rangle \in S\), \(A_i \cap A_j \in \Sigma\),

\[
p_i = \mu(A_i) \quad \text{and} \quad p_{ij} = \mu(A_i \cap A_j).
\]

The form of the definition is exactly the same as before (the sense 1 case) except for the weakened requirement on \(\Sigma\). A hidden variable theory in sense 2 returns the probabilities \(p_i\) and \(p_{ij}\) where \(\langle i, j \rangle \in S\), but need not assign a probability to \((A_k \cap A_l)\) if \(\langle k, l \rangle \in S\). The question to ask in parallel to Pitowsky’s is this.

**Question:** Under what conditions does \(\langle p_1, ..., p_n, ... p_{ij} ... \rangle\) admit a hidden variable theory in sense 2?

I do not have a general answer. It is not at all clear to me that there is a neat and simple general answer. (This might be a good research problem to pursue.) But it is not difficult to verify the that the condition of admitting a hidden variable theory in sense 2 is weaker than the counterpart condition involving sense 1. In particular, the experimental probability data set

\[
\langle p_1, p_2, p_3, p_4, p_{13}, p_{14}, p_{23}, p_{24} \rangle = \langle 1/2, 1/2, 1/2, 1/2, 1/2, 1/2, 1/2, 1/2 \rangle
\]

of Clauser-Horne type (cited above) admits a hidden variable theory in sense 2 (though not in sense 1). Here is the argument.
Proof. Let $X$ be the set of all (sixteen) ordered quadruples of 0 and 1. Let

- $A_1 = \text{the set of elements of } X \text{ of form } 1_{---}$
- $A_2 = \text{the set of elements of } X \text{ of form } 1_{--}$
- $A_3 = \text{the set of elements of } X \text{ of form } --1$
- $A_4 = \text{the set of elements of } X \text{ of form } ---1$

Let $\Sigma$ consist of the empty set and all disjoint unions of the following (4 element) sets:

- $(A_1 \cap A_3)$
- $(A_1 - A_3)$
- $(A_3 - A_1)$
- $(A_1 \cup A_3)^c$
- $(A_1 \cap A_4)$
- $(A_1 - A_4)$
- $(A_4 - A_1)$
- $(A_1 \cup A_4)^c$
- $(A_2 \cap A_3)$
- $(A_2 - A_3)$
- $(A_3 - A_2)$
- $(A_2 \cap A_3)^c$
- $(A_2 \cap A_4)$
- $(A_2 - A_4)$
- $(A_4 - A_2)$
- $(A_2 \cap A_4)^c$

(Here, for any subset $A$ of $X$, $A^c$ is the complement set $X - A$.)

The measure $\mu : \Sigma \to [0, 1]$ is defined by making the following explicit assignments to the sets in the display, and then extending to all other sets in $\Sigma$ using finite additivity.

$$
\begin{array}{cccc}
3/8 & 3/8 & 0 & 3/8 \\
1/8 & 1/8 & 1/2 & 1/8 \\
1/8 & 1/8 & 1/2 & 1/8 \\
3/8 & 3/8 & 0 & 3/8 \\
\end{array}
$$

(Note that these are the same entries that appeared in the table in section 3. But here they have a different significance.)

We claim that $\Sigma$ satisfies conditions (F1), (F2), and (F3*). The four sets in each column of the first display are disjoint. And the union of the four, in each case, is $X$. So $X \in \Sigma$. This gives us (F1). (F3*) clearly holds by the very way $\Sigma$ was defined. So the only thing we need to check is that $\Sigma$ is closed under complementation. It suffices to consider a few representative cases.

First, let $A$ be a set in $\Sigma$ that is the union of one, two, or three sets in the same column. Then $A^c$ is the union of the remaining sets in that column, and so belongs to $\Sigma$. For example,

$$[(A_1 \cap A_3) \cup (A_3 - A_1)]^c = (A_1 - A_3) \cup (A_1 \cup A_3)^c.$$

And if $A$ the union of all four sets in the same column, then $A = X$, and $A^c$ is the empty set. So in this case too, $A^c$ belongs to $\Sigma$.

Next, consider the case where $A$ is a disjoint union of basic sets coming from more than one column. We claim that no more than two columns can be involved. So, for example, suppose $A$ is a disjoint union of $(A_1 \cap A_3)$, $(A_4 - A_1)$, and some third set $B$. (Note that the first two are disjoint, because $(A_1 \cap A_3)$ is a subset of $A_1$, and $(A_4 - A_1)$ is a subset of the complement of $A_1$.) There are only a few possibilities for $B$. The only basic sets disjoint from $(A_1 \cap A_3)$ are:
\[(A_1 - A_3) \quad (A_2 - A_3)\]
\[(A_3 - A_1) \quad (A_4 - A_1)\]
\[(A_1 \cup A_3)^c \quad (A_1 \cup A_4)^c \quad (A_2 \cup A_3)^c\]

And of these, the only ones disjoint from \((A_4 - A_1)\) are

\[(A_1 - A_3)\]
\[(A_1 \cup A_4)^c\]

So, the only disjoint unions of basic sets containing \((A_1 \cap A_3)\) and \((A_4 - A_1)\) as subsets are:

\[(A_1 \cap A_3) \cup (A_4 - A_1)\]
\[(A_1 \cap A_3) \cup (A_4 - A_1) \cup (A_1 - A_3)\]
\[(A_1 \cap A_3) \cup (A_4 - A_1) \cup (A_1 \cup A_4)^c\]
\[(A_1 \cap A_3) \cup (A_4 - A_1) \cup (A_1 - A_3) \cup (A_1 \cup A_4)^c\].

(Indeed, by systematically interchanging the roles of \(A_1\) and \(A_2\), \(A_3\) and \(A_4\), and for all \(i\), \(A_i\) and \((A_i)^c\), we can reduce to one of these forms any disjoint union of basic sets coming from more than one column.)

It is easy to check that the respective complements of these four are all expressible as unions of disjoint unions of basic sets, and so belong to \(\Sigma\):

\[
\left[(A_1 \cap A_3) \cup (A_4 - A_1)^c\right]^c = (A_1 - A_3) \cup (A_1 \cup A_4)^c
\]
\[
\left[(A_1 \cap A_3) \cup (A_4 - A_1)^c \cup (A_1 - A_3)^c\right]^c = (A_1 \cup A_4)^c
\]
\[
\left[(A_1 \cap A_3) \cup (A_4 - A_1)^c \cup (A_1 \cup A_4)^c\right]^c = (A_1 - A_3)
\]
\[
\left[(A_1 \cap A_3) \cup (A_4 - A_1)^c \cup (A_1 - A_3) \cup (A_1 \cup A_4)^c\right]^c = \emptyset.
\]

This completes the argument that \(\Sigma\) satisfies conditions (F1), (F2), and (F3*). It also follows easily that \(\mu\) satisfies (M1) and (M2). The latter additivity requirement was, in effect, built into the definition of \(\mu\).

And the former normalization condition follows from additivity and the fact that the disjoint union of the four basic sets in any column is \(X\), e.g.,

\[
\mu(X) = \mu[(A_1 \cap A_3) \cup (A_1 - A_3) \cup (A_3 - A_1) \cup (A_1 \cup A_3)^c]
\]
\[
= \frac{3}{8} + \frac{1}{8} + \frac{1}{8} + \frac{3}{8} = 1.
\]

So we may conclude that \((X, \Sigma, \mu)\) qualifies as a generalized probability space. It only remains to show that it returns the probabilities with which we began. But this is immediate. We see from the entries in
the top row of the numerical array that
\[
\mu(A_1 \cap A_3) = \frac{3}{8} = p_{13} \\
\mu(A_1 \cap A_4) = \frac{3}{8} = p_{14} \\
\mu(A_2 \cap A_3) = 0 = p_{23} \\
\mu(A_2 \cap A_4) = \frac{3}{8} = p_{24}
\]

And, clearly, we also have
\[
\mu(A_1) = \mu(A_1 \cap A_3) + \mu(A_1 - A_3) = \frac{3}{8} + \frac{1}{8} = \frac{1}{2} = p_1,
\]
and, similarly,
\[
\mu(A_2) = \frac{1}{2} = p_2, \\
\mu(A_3) = \frac{1}{2} = p_3, \\
\mu(A_4) = \frac{1}{2} = p_4.
\]
Thus, as claimed, our data set admits a hidden variable theory in sense 2.

It will be instructive now to verify directly that it is not possible to extend the generalized probability space \((X, \Sigma, \mu)\) to a (full) classical probability space \((X, \Sigma', \mu')\). (We know in advance that it is not possible, since our data set does not admit a hidden variable theory in sense 1.)

Suppose there is a classical probability space \((X, \Sigma', \mu')\) built over the same underlying set \(X\) such that (i) \(\Sigma \subseteq \Sigma'\) and (ii) the restriction of \(\mu'\) to \(\Sigma\) is \(\mu\), i.e., \(\mu'(A) = \mu(A)\) for all sets \(A\) in \(\Sigma\). We proceed to derive a contradiction.

First we claim that every singleton subset of \(X\) belongs to \(\Sigma'\). Consider, for example, the subset \(\{0, 1, 0, 1\}\). It can be expressed as:
\[
\{0, 1, 0, 1\} = (A_2 \cap A_4) \cap (A_1 \cap A_3)^c.
\]
Since \((A_2 \cap A_4)\) and \((A_1 \cap A_3)^c\) both belong to \(\Sigma'\), and since \(\Sigma'\) is closed under intersection (condition (F4)), it follows immediately that \(\{0, 1, 0, 1\}\) belongs to \(\Sigma'\). The other cases are handled similarly. (Of course, \((A_2 \cap A_4)\) and \((A_1 \cap A_3)^c\) also belong to \(\Sigma\) but we cannot conclude that \(\{0, 1, 0, 1\}\) belongs to it, because \(\Sigma\) is not closed under arbitrary intersection.)

To simplify notation, let’s write, for example, \(\mu'(0, 1, 0, 1)\) for \(\mu'(\{0, 1, 0, 1\})\).

Next, note that since \(\mu'(A_2 \cap A_3) = 0\) (and since \(A \subseteq B \Rightarrow \mu'(A) \leq \mu'(B)\) for all sets \(A, B\) in \(\Sigma'\)), we have:
\[
\mu'(0, 1, 1, 0) = \mu'(0, 1, 1, 1) = \mu'(1, 1, 1, 0) = \mu'(1, 1, 1, 1) = 0.
\]
Similarly, since $\mu^\prime(\overline{A_2 \cup A_3}) = 0$,

$$\mu^\prime(0, 0, 0, 0) = \mu^\prime(0, 0, 0, 1) = \mu^\prime(1, 0, 0, 0) = \mu^\prime(1, 0, 0, 1) = 0.$$  

Now, consider the numbers

$$x = \mu^\prime(0, 1, 0, 0)$$
$$y = \mu^\prime(0, 1, 0, 1)$$
$$z = \mu^\prime(1, 1, 0, 0)$$
$$w = \mu^\prime(0, 0, 1, 1).$$

We claim that it follows from what we have so far that:

(i) $x + z = \frac{1}{8}$
(ii) $x + y = \frac{3}{8}$
(iii) $y + w = \frac{1}{8}$

For (i), note that

$$\frac{1}{8} = \mu^\prime(A_2 - A_4) = \mu^\prime(\{0, 1, 0, 0\}, \{0, 1, 1, 0\}, \{1, 1, 0, 0\}, \{1, 1, 1, 0\}))$$
$$= x + 0 + z + 0.$$

Similarly, we have

$$\frac{3}{8} = \mu^\prime(A_1 \cup A_3)^c = \mu^\prime(\{0, 0, 0, 0\}, \{0, 0, 0, 1\}, \{0, 1, 0, 0\}, \{0, 1, 0, 1\}))$$
$$= 0 + 0 + x + y.$$  

and

$$\frac{1}{8} = \mu^\prime(A_4 - A_1) = \mu^\prime(\{0, 0, 0, 1\}, \{0, 0, 1, 1\}, \{0, 1, 0, 1\}, \{0, 1, 1, 1\}))$$
$$= x + w + y + 0.$$

By (i), now, $\frac{1}{8} - x = z \geq 0$. So $x \geq \frac{1}{8}$. But by (iii) and (ii),

$$0 \leq w = \frac{1}{8} - y = \frac{1}{8} - \left(\frac{3}{8} - x\right) = -\frac{1}{4} + x,$$

and, so, $\frac{1}{4} \leq x$. Thus we may conclude that $\frac{1}{4} \leq x \leq \frac{1}{8}$, which is absurd.
References


